

# A high-level overview of Bayesian Statistics

Bayesian Idea: Treat a parameter as having an unknown Distribution to be estimated, rather than as an unknown fixed number to be estimated.

Ex 1 Predict end-of-season win percentage from mid-season Wins and Losses.

Beta-Binomial model

$$\begin{cases} W \sim \text{Binomial}(n, p) \\ p \sim \text{Beta}(\alpha, \beta) \end{cases}$$

When we model the latent team's win probability  $p$  using a Beta Prior, we encode the prior information that  $p$  is more likely to be near  $1/2$  (say, in  $.3, .7$ ) than to be very near 0 or 1.

Then, we found using Bayes Rule that the posterior distribution  $p|W, L$  is

$$p|W, L \sim \text{Binomial}(n + \alpha + \beta - 2, \frac{W + \alpha - 1}{W + L + \alpha + \beta - 2})$$

and the Bayes estimate is the posterior mean

$$\hat{p}^{(Bayes)} = \mathbb{E}[p|W, L] = \frac{W + (\alpha - 1)}{W + L + (\alpha + \beta - 2)}$$

Ex 2 Predict end-of-season batting average from mid-season batting average and number of at-bats.

Normal-Normal Model

$i$  = player  $i$   
 $H_i$  = # hits,  $N_i$  = # at-bats,  $X_i = \frac{H_i}{N_i}$

$$X_i \sim N(\mu_i, \sigma_i^2)$$

$$\sigma_i^2 = C/N_i, \quad C \text{ known}$$

$$\mu_i \sim N(\mu, \tau^2)$$

When we model player  $i$ 's latent quality  $\mu_i$  using a **Normal Prior**, we encode the prior information that player  $i$  is also a baseball player, allowing us to share strength across players.

Then, we found using Bayes Rule that the **posterior distribution**  $\mu_i | X$  is

$$\mu_i | X \sim N\left(\frac{\frac{X_i \tau^2 + \frac{\mu}{C}}{\frac{1}{\sigma_i^2} + \frac{1}{C}}}{\frac{1}{\sigma_i^2} + \frac{1}{C}}, \frac{1}{\frac{1}{\sigma_i^2} + \frac{1}{C}}\right)$$

and the Bayes estimate is the **posterior mean**

$$\hat{\mu}_i^{(\text{Bayes})} = \frac{X_i \tau^2 + \frac{\mu}{C}}{\frac{1}{\sigma_i^2} + \frac{1}{C}}$$

Using  $\hat{\mu}$  and  $\hat{\tau}^2$  since  $\mu, \tau$  are unknown.

### Ex 3

## Bayesian Regression

Suppose we have a regression model

$$y = X\beta + \varepsilon \quad \text{with mean zero noise } \mathbb{E}\varepsilon = 0.$$

For instance, our Park effects model (so we want to estimate the park effects  $\beta$ ), or a Power Score model like Bradley Terry.

A Bayesian puts a prior distribution on the parameters  $\beta$  (and, a distribution on the error term  $\varepsilon$ ) to end up with a posterior distribution  $\beta|X$  or a posterior mean  $\mathbb{E}[\beta|X]$ , quantifying our best guess of  $\beta$  after seeing the data  $X$ , while also incorporating our prior information.

\* "Standard" Regression Setup:

$$y = X\beta + \varepsilon$$

$$\left. \begin{array}{l} \text{Homoscedasticity} \rightarrow \varepsilon_i \text{ iid } \mathcal{N}(0, \sigma^2) \\ \Rightarrow \varepsilon \sim \mathcal{N}(0, \sigma^2 I) \end{array} \right\} \Rightarrow y \sim \mathcal{N}(X\beta, \sigma^2 I)$$

\* Bayesian Regression ex1: uninformative Uniform Prior  $\beta \propto 1$

\* Bayesian Regression ex2: Normal Prior  $\beta \sim \mathcal{N}(0, \lambda)$

## Bayesian Regression v1

$$\left\{ \begin{array}{l} y \sim N(X\beta, \sigma^2 I) \\ \beta \propto 1 \text{ uniform prior} \\ \text{Find posterior dist } \beta|X \end{array} \right.$$

$$P(\beta|y) = \frac{P(y|\beta)P(\beta)}{P(y)} \quad \text{by Bayes Rule}$$

$$\propto P(y|\beta) \cdot P(\beta) \quad \text{since } P(y) \text{ doesn't depend on } \beta$$

$$\propto P(y|\beta) \quad \text{since } P(\beta) \propto 1$$

$$\propto \prod_{i=1}^n \exp\left(-\frac{1}{2\sigma^2} (y_i - x_i^T \beta)^2\right)$$

$$= \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - x_i^T \beta)^2\right)$$

Bayes estimate of  $\beta$  is to choose the  $\beta$  which maximizes the posterior probability

Maximum A-posteriori (MAP):

$$\hat{\beta}^{(\text{Bayes})} = \underset{\beta}{\text{argmax}} P(\beta|y)$$

$$= \underset{\beta}{\text{argmax}} \sum_{i=1}^n (y_i - x_i^T \beta)^2 = \hat{\beta}^{(\text{OLS})} = (X^T X)^{-1} X^T y$$

\* With an uninformative uniform prior, Bayesian Regression is the same as Ordinary least squares!

## Bayesian Regression v2

$$\begin{cases} y \sim N(X\beta, \sigma^2 I) \\ \beta \sim N(0, \frac{1}{\lambda} I) \\ \text{Find posterior dist } \beta | X \end{cases}$$

$$P(\beta | y) = \frac{P(y | \beta) P(\beta)}{P(y)} \quad \text{by Bayes Rule}$$

$$\propto P(y | \beta) \cdot P(\beta) \quad \text{since } P(y) \text{ doesn't depend on } \beta$$

$$= P(N(X\beta, \sigma^2 I) = y) \cdot P(\beta = N(0, \lambda))$$

$$\propto \prod_{i=1}^n \exp\left(-\frac{1}{2} \frac{(x_i^T \beta - y_i)^2}{\sigma^2}\right) \cdot \prod_{j=1}^p \exp\left(-\frac{1}{2} \frac{(\beta_j - 0)^2}{1/\lambda}\right)$$

$$= \exp\left(-\frac{1}{2} \left[ \lambda \sum_{j=1}^p \beta_j^2 + \sum_{i=1}^n \frac{(x_i^T \beta - y_i)^2}{\sigma^2} \right]\right)$$

$$= \exp\left(-\frac{1}{2\sigma^2} \left[ \lambda \sigma^2 \sum_{j=1}^p \beta_j^2 + \sum_{i=1}^n (y_i - x_i^T \beta)^2 \right]\right)$$

Bayes estimate of  $\beta$  is to choose the  $\beta$  which maximizes the posterior probability **Maximum A-posteriori (MAP)**:

$$\hat{\beta}^{(\text{Bayes})} = \underset{\beta}{\operatorname{argmax}} P(\beta | y) = \underset{\beta}{\operatorname{argmax}} \lambda \sum_{j=1}^p \beta_j^2 + \sum_{i=1}^n (y_i - x_i^T \beta)^2$$

$$\left(\text{letting } \lambda \leftarrow \frac{\lambda}{\sigma^2}\right) = \hat{\beta}^{(\text{Ridge})} = (X^T X + \lambda I)^{-1} X^T y.$$

\* With a Normal prior, Bayesian Regression is Ridge Regression. This makes perfect sense:  $\beta \sim N(0, \lambda I)$  encourages  $\beta$  to be closer to 0.  $\lambda$  controls by how much, and that is the same as penalizing  $\beta$ !

\* So far we have mainly focused on finding the **posterior mean**, i.e. our best estimate of the parameters after seeing the data.

\* To **make decisions** in sports (e.g. player valuation or play selection), we need not only know our best estimate of the value of the player/play/decision, but also **uncertainty quantification** (e.g., error bars or prediction intervals) to describe how confident/certain we are about a value.

Ex Suppose we have a prediction of end-of-season win percentage to be 74%.

A 95% prediction interval of [72%, 76%] is much different from [60%, 88%] and from [0%, 100%].

Bayesian Idea Estimate the FULL Posterior Distribution of an unknown parameter. This gets us error bars (uncertainty quantification), with which we can create more complete decisions.

\* In a general Bayesian statistical model, the parameters don't all have Gaussian priors or likelihoods, and we can't write on paper a closed-form analytical solution for the posterior distribution.

\* Typically need to approximate the posterior distribution using MCMC sampling methods like

- Gibbs sampling
- Hamiltonian Monte Carlo
- No U-Turn Sampling,

which we won't cover here (take Shane's class), using a probabilistic programming language [ Stan, Jags, NumPyro ]

\* Next Time : a full example of Bayesian modeling in sports.

Takeaway Create a fully Bayesian probability model, which allows quantify both the effect size and uncertainty of a phenomenon of interest by approximating the full posterior distribution.