

Probability Primer

Agenda Discrete uniform distribution, Expected Value E , Bernoulli, Binomial, linearity of expectation, Independence, Variance, Normal distribution, continuous probability, density, cdf, conditional probability, law of total probability, Bayes Rule, Poisson

HW { • watch Wyner's probability lectures from Moneyball

Dice Let X represent the roll of a dice,

$$X = \begin{cases} 1 & \text{with probability } 1/6 \\ 2 & \text{up} & 1/6 \\ 3 & & \\ 4 & \dots & 1/6 \\ 5 & & \\ 6 & & \end{cases}$$

Formally, $X \sim \text{Uniform}(\{1, 2, 3, 4, 5, 6\})$.

X is a discrete random variable discrete because the set of outcomes $\Omega = \{1, 2, 3, 4, 5, 6\}$ is finite (countable).

Q What is the average value of a dice roll?

$$\frac{1+2+3+4+5+6}{6} = 3.5$$

But why is 3.5 the average value?

One nice interpretation is this, if you roll a dice $n = 10$ trillion times, you'd expect

just about 10 trillion 1's, 10 trillion 2's, etc.
and the numeric average of these values
will be just about 3.5 (the frequency argument).

* What if now we define a new die Y

as the random variable $Y = \begin{cases} 1 & \text{up } 0 \\ 2 & \text{up } 2/6 \\ 3 & \\ 4 & \rightarrow \text{up } 1/6 \\ 5 & \\ 6 & \end{cases}$

What then is the value of an average die roll?
Again consider the frequency argument.

If we roll $n = 60$ trillion dice, then

10 trillion each will be 3, 4, 5, 6

20 trillion will be 2

so the sum of all the dice rolls is

$$\frac{2(20 \text{ trillion}) + 3(10 \text{ trillion}) + 4(10 \text{ trillion}) + 5(10 \text{ trillion}) + 6(10 \text{ trillion})}{20 \text{ trillion} + 10 \text{ trillion} + \dots + 10 \text{ trillion}}$$

$$= \frac{2 \cdot 2 + 3 + 4 + 5 + 6}{6} = \frac{22}{6} = 3.67$$

* Written in another way,

$$2\left(\frac{2}{6}\right) + 3\left(\frac{1}{6}\right) + 4\left(\frac{1}{6}\right) + 5\left(\frac{1}{6}\right) + 6\left(\frac{1}{6}\right)$$

$$= \frac{2 \cdot 2 + 3 + 4 + 5 + 6}{6} = 3.67$$

So, simply multiply each outcome by its probability and then sum!

Def The expected value of a discrete random variable X is

$$\mathbb{E}X = \sum_x x P(X=x)$$

- discrete random variable X means X can be one of countably many outcomes Ω
e.g. Dice, $\Omega = \{1, 2, 3, 4, 5, 6\}$
- The summation \sum_x is over the outcomes $x \in \Omega$
- $P(X=x)$ is the probability that the random variable X equals the value x
e.g. Dice, $x=5$, $P(X=5) = 1/6$

Coin Flip

Let X represent a coin flip,

$$X = \begin{cases} 1 & \text{with probability } p \text{ (heads)} \\ 0 & \text{w.p. } 1-p \text{ (tails)} \end{cases}$$

Formally,

$$X \sim \text{Bernoulli}(p)$$

p here is a parameter, $p \in [0, 1]$.

The expected value of a $\text{Bernoulli}(p)$ random variable, or informally the average value of a biased coin flip, is

$$\begin{aligned} \mathbb{E}X &= \sum_x x \mathbb{P}(X=x) \\ &= \sum_{x \in \{0,1\}} x \mathbb{P}(X=x) \\ &= 0 \cdot \mathbb{P}(X=0) + 1 \cdot \mathbb{P}(X=1) \\ &= \mathbb{P}(X=1) \\ &= p. \end{aligned}$$

Shaqtin' a fool

A free throw shooter named _____ takes n free throw shots.

He makes a free throw with prob. p .

The random variable X representing the number of shots he makes is

$$X = \sum_{i=1}^n X_i$$

$X_i \stackrel{iid}{\sim} \text{Bernoulli}(p)$ iid = independent and identically distributed

i.e. $X_i = \mathbb{1}_{\left\{ \begin{array}{l} \text{ith shot} \\ \text{is made} \end{array} \right\}} = \begin{cases} 1 \text{ w.p. } p \\ 0 \text{ w.p. } 1-p \end{cases}$

X is the sum of n Bernoulli(p) R.V.'s,

$$X \sim \text{Binomial}(n, p)$$

How many shots will he make on average?

Well, the expected value of a Binomial R.V. is

$$EX = \sum_x x P(X=x) = \sum_{x=0}^n x P(X=x)$$

and $P(X=x) = P(\text{make } x \text{ of } n \text{ shots, } x \in \{0, \dots, n\})$

for one ordering, e.g.

$$\underbrace{\frac{1}{\quad} \frac{1}{\quad} \frac{1}{\quad} \cdots \frac{1}{\quad}}_{x \text{ made shots}} \quad \underbrace{\frac{0}{\quad} \frac{0}{\quad} \cdots \frac{0}{\quad}}_{n-x \text{ missed shots}}$$

the probability that this exact ordering occurs is

$$\underbrace{p \cdot p \cdot p \cdots p}_{x \text{ made shots}} \cdot \underbrace{(1-p)(1-p) \cdots (1-p)}_{n-x \text{ missed shots}}$$

$$= p^x (1-p)^{n-x}$$

and the number of such orderings with x 1's is

$$\binom{n}{x} = n \text{ choose } x = \frac{n!}{x! (n-x)!}$$

$$\text{so } P(X=x) = \binom{n}{x} p^x (1-p)^{n-x}$$

$$\text{so } EX = \sum_x x P(X=x) = \sum_{x=0}^n \binom{n}{x} p^x (1-p)^{n-x}$$

How the fuck do we sum this though?

The Binomial Theorem goes through this, using induction. But there's an easier way!

Theorem (Linearity of Expectation)

$$\mathbb{E}\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n \mathbb{E}[X_i]$$

and $\mathbb{E}[cX] = c\mathbb{E}[X]$, $c \in \mathbb{R}$

Idea Average value of the sum of 2 dice rolls is the sum of the averages! Think about it, using the frequency argument!

Proof Idea

$$\mathbb{E}[X+Y] = \sum_x \sum_y (x+y) P(X=x, Y=y)$$

↳ and

$$= \sum_x x \sum_y P(X=x, Y=y) + \sum_y y \sum_x P(X=x, Y=y)$$

$$= \sum_x x P(X=x) + \sum_y y P(Y=y)$$

$$= \mathbb{E}X + \mathbb{E}Y.$$

$$\begin{aligned}
 \text{and } E[cX] &= \sum_{cx} cx P(cX=cx) \\
 &= \sum_x cx P(X=x) \\
 &= c \sum_x x P(X=x) \\
 &= c EX
 \end{aligned}$$

Therefore, when $X \sim \text{Binomial}(n, p)$,

$$X = \sum_{i=1}^n X_i \quad \text{where } X_i \stackrel{iid}{\sim} \text{Bernoulli}(p)$$

$$EX = \sum_{i=1}^n EX_i = \sum_{i=1}^n p = np. \quad \text{Easy!}$$

So, if _____ takes n shots, then he will on average make np of them!

Now, recall for $X \sim \text{Binomial}(n, p)$, the number of made shots, that $X = \sum_{i=1}^n X_i$ where $X_i \stackrel{iid}{\sim} \text{Bernoulli}(p)$.

iid means independent and identically distributed. identically distributed here means each X_i is similarly a

Bernoulli (p) random variable.

But what does independence mean formally?

def X and Y are independent random variables means

$$P(X=x, Y=y) = P(X=x) \cdot P(Y=y)$$

* Why is this definition a good one?

Frequently argument:

1 trillion pairs of independent shots.

$\approx p$ trillion made 1st shots

- of these, $\approx p$ fraction are paired with a 2nd shot make
- of these, $\approx 1-p$ fraction are paired with a 2nd shot miss

$\approx 1-p$ trillion missed 1st shots

- $\approx p$
- $\approx 1-p$



$\approx p^2$ trillion pairs (made, made)

$\approx p \cdot (1-p)$ trillion pairs (made, missed)

$\approx (1-p) \cdot p$ trillion pairs (missed, made)

$\approx (1-p)^2$ trillion pairs (missed, missed)

□

Thm If X and Y are independent then

$$\mathbb{E}(XY) = (\mathbb{E}X) \cdot (\mathbb{E}Y)$$

Proof Assume $P(X=x, Y=y) = P(X=x) \cdot P(Y=y)$.

Then

$$\begin{aligned}\mathbb{E}(X \cdot Y) &= \sum_x \sum_y xy P(X=x, Y=y) \\ &= \sum_x \sum_y xy P(X=x) \cdot P(Y=y) \\ &= \sum_x x P(X=x) \cdot \sum_y y P(Y=y) \\ &= (\mathbb{E}X) \cdot (\mathbb{E}Y).\end{aligned}$$

Recall if $X \sim \text{Binomial}(n, p)$ then $\mathbb{E}X = np$,
so _____ will make np free throws on average.

But what will the spread or deviation
from this spread look like ??

Def The variance of a random variable X is
the expected squared deviation from its mean,
$$\text{VAR}(X) = \mathbb{E}[(X - \mathbb{E}X)^2]$$

Large variance = on average, X is far from its mean, since $(X - \mathbb{E}X)^2$ is large often.

Theorem (Funny Variance Formula)

$$\text{var}(X) = \mathbb{E}[X^2] - (\mathbb{E}X)^2$$

Proof $\text{var}(X) = \mathbb{E}(X - \mathbb{E}X)^2$

$$= \mathbb{E}(X^2 - 2X\mathbb{E}X + (\mathbb{E}X)^2)$$

$$= \mathbb{E}X^2 - \mathbb{E}[2X\mathbb{E}X] + \mathbb{E}[(\mathbb{E}X)^2]$$

by linearity of expectation

$$= \mathbb{E}X^2 - 2\mathbb{E}X \cdot \mathbb{E}X + (\mathbb{E}X)^2$$

by linearity of expectation

because $\mathbb{E}X \in \mathbb{R}$, $(\mathbb{E}X)^2 \in \mathbb{R}$
are real numbers

$$= \mathbb{E}X^2 - (\mathbb{E}X)^2$$

So, the variance of _____'s number of made free throws is

$$\begin{aligned} \text{VAR}(X) &= \mathbb{E}X^2 - (\mathbb{E}X)^2 \\ &= \mathbb{E}X^2 - (np)^2 \end{aligned}$$

$$\text{and } \mathbb{E}X^2 = \mathbb{E}\left(\sum_{i=1}^n X_i\right)^2 \quad X_i \stackrel{\text{iid}}{\sim} \text{Ber}(p)$$

$$= \mathbb{E}\left(\sum_{i=1}^n \sum_{j=1}^n X_i X_j\right)$$

$$= \mathbb{E}\left(\sum_{i \neq j} X_i X_j + \sum_{i=1}^n X_i\right)$$

$$= \sum_{i \neq j} \mathbb{E}(X_i X_j) + \sum_{i=1}^n \mathbb{E}X_i$$

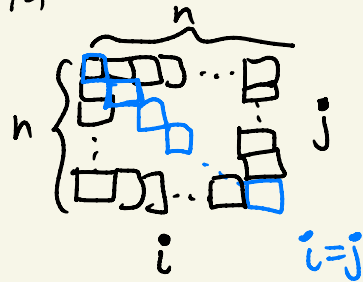
by linearity of expectation

$$= \sum_{i \neq j} (\mathbb{E}X_i) \cdot (\mathbb{E}X_j) + \sum_{i=1}^n \mathbb{E}X_i$$

by independence

$$= \sum_{i \neq j} p \cdot p + \sum_{i=1}^n p$$

$$= (n^2 - n) \cdot p^2 + n \cdot p$$



$$\begin{aligned} \text{Then } \text{VAR}(X) &= n^2 p^2 - np^2 + np - n^2 p^2 = n(p - p^2) \\ &= np(1 - p). \end{aligned}$$

The standard deviation of a R.V. X is the square root of the variance,

$$sd(X) = \sqrt{\text{var}(X)}$$

Because $\text{var}(X) = \mathbb{E}(X - \mathbb{E}X)^2$ is on the squared scale, the s.d. uses a more grounded unit of measurement.

$$X \sim \text{Binomial}(n, p)$$

$$X = \sum_{i=1}^n X_i, \quad X_i \stackrel{\text{iid}}{\sim} \text{Ber}(p)$$

$$\mathbb{E}X = np, \quad sd(X) = \sqrt{np(1-p)}$$

Central limit theorem the (Rescaled) sum of iid random variables X_i with mean $\mu = \mathbb{E}X_i$ and s.d. $\sigma = \sqrt{\text{var}(X_i)}$, $S_n = \sum_{i=1}^n X_i$, converges to a standard Normal (Gaussian) Random Variable $Z \sim \mathcal{N}(0, 1)$,

$$\lim_{n \rightarrow \infty} \mathbb{P}\left(a \leq \frac{S_n - n\mu}{\sigma\sqrt{n}} \leq b\right) = \mathbb{P}(a \leq Z \leq b)$$

$X \sim \text{Binomial}(n, p)$ is $S'_n = X$

$$\frac{S'_n - n\mu}{\sigma\sqrt{n}} \approx \mathcal{N}(0, 1) \text{ for } n \text{ large}$$

$$S'_n \approx \mathcal{N}(n\mu, \sqrt{n}\sigma)$$

$$S'_n = X, \quad \mu = p, \quad \sigma = \sqrt{p(1-p)}$$

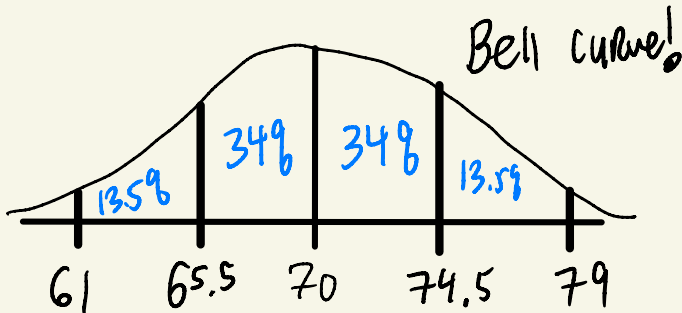
So, $X \approx \mathcal{N}(np, \sqrt{np(1-p)})$.

_____ takes $n = 100$ shots.

Assume we know his $p = .7$

$$\sqrt{np(1-p)} = 10\sqrt{(0.7)(0.3)} = 10\sqrt{0.21} = 10(0.458) \approx 4.5$$

$$np = 100(.7) = 70$$



Normal
dist
area
values.

* But what really is the normal distribution?

Discrete probability: $P(X=x)$ defines a R.V.

Now, if $X \sim N(\mu, \sigma^2)$, X can take on any real number in \mathbb{R} , so if $P(X=x) > 0$ for each $x \in \mathbb{R}$ then $\sum_{x \in \mathbb{R}} P(X=x) = \infty \neq 1$, a contradiction...

So, in continuous probability, where a Random variable X can take on any of uncountably many values in \mathbb{R} , X is defined

by its density function $f_X(x) = P(X \in [x, x+dx])$

I hope you like calculus!

Let's start with the

CDF (cumulative distribution function)

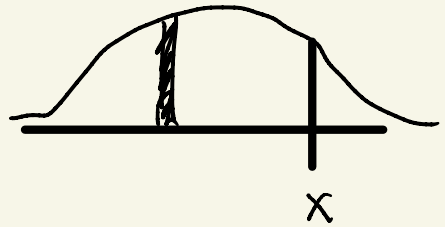
$$F_X(x) = P(X \leq x)$$

which is easy to understand.

$$\text{Well, } F(x) = P(X \leq x)$$

$$= \int_{-\infty}^x P(X \in [x, x+dx]) dx$$

by def. of integral as area under curve



$$= \int_{-\infty}^x f_x(x) dx$$

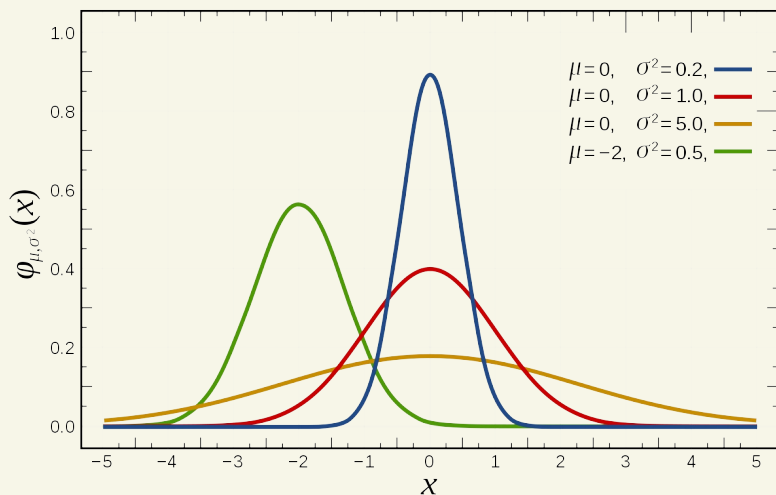
So we define the density as the

derivative of the cdf,

$$f_x(x) = F'_x(x)$$

Def The normal distribution with mean μ and variance σ^2 , written $X \sim \mathcal{N}(\mu, \sigma^2)$,

has density $f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$



Before, in discrete probability, $\mathbb{E}X = \sum_x x \mathbb{P}(X=x)$

Now, in continuous probability, $\mathbb{P}(X=x)$ does not exist, and it is replaced by $f(x) = \mathbb{P}(X \in (x, x+dx))$.

The analog of \sum_x in continuous space is \int_x because we sum up infinitesimally small rectangles.

Def The expectation of a continuous R.V. X is

$$\mathbb{E}X = \int x f(x) dx$$

and of a function of X is

$$\mathbb{E}g(X) = \int g(x) f(x) dx.$$

If $X \sim N(\mu, \sigma^2)$ then

$$\mathbb{E}X = \int x f(x) dx \stackrel{\text{HW}}{=} \mu \quad \text{and}$$

$$\mathbb{E}X^2 = \int x^2 f(x) dx \stackrel{\text{HW}}{=} \sigma^2 \Rightarrow \text{Var}(X) = \sigma^2.$$

The standard normal distribution $Z \sim N(0, 1)$

has density $\phi(x) := \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$

and std. normal CDF $\Phi(x) := \int_{-\infty}^x \phi(x) dx$

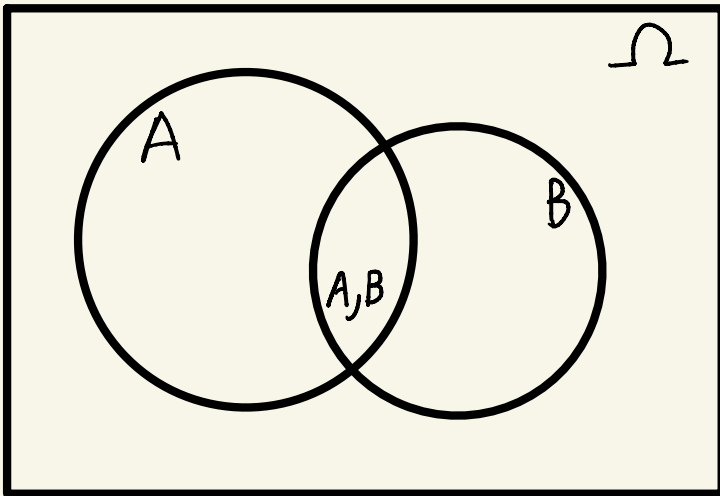
which does not have a closed form.

def the conditional probability of $A|B$ "A given B" is defined by

$$P(A|B) := \frac{P(A, B)}{P(B)}$$

A, B "A and B"

To understand why this is a good definition, see Venn diagram:



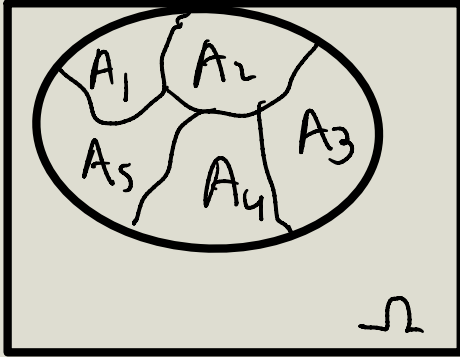
Ω = entire sample space (set of outcomes)

Given we're in B, the total prob. is $P(B)$,
and the prob. of hitting A is $P(A, B)$,
so $P(A|B) := P(A, B) / P(B)$.

Law of Disjoint Unions

If A_1, \dots, A_n partitions the set A ,

so $A_1 \cup A_2 \cup \dots \cup A_n = A$ and $A_i \cap A_j = \emptyset \quad \forall i \neq j$



then

$$P(A) = \sum_{i=1}^n P(A_i)$$

Thm If X, Y are independent then $P(X|Y) = P(X)$

$$\text{Pf} \quad P(X|Y) = \frac{P(X, Y)}{P(Y)} = \frac{P(X) \cdot P(Y)}{P(Y)} = P(X).$$

$$P(A|B) := \frac{P(A, B)}{P(B)}$$

and so we also have $P(B|A) = \frac{P(B, A)}{P(A)}$

$$\Rightarrow P(A|B) = \frac{P(A, B)}{P(B)} = \frac{P(B, A)}{P(B)} = \frac{P(B|A)P(A)}{P(B)}$$

Bayes' Rule $P(A|B) = \frac{P(B|A)P(A)}{P(B)}$

* Extremely important!

Ex Suppose half of USA is male and half is female,

But $\frac{1}{3}$ of NBA fans are female,

Also, $\frac{1}{3}$ of Americans are NBA fans.

What's the probability of a male/female being an NBA fan?

$$P(M) = \frac{1}{2}, \quad P(F) = \frac{1}{2}, \quad P(M/NBA) = \frac{2}{3}, \quad P(F/NBA) = \frac{1}{3}$$

$$P(NBA|M) = \frac{P(M/NBA)P(NBA)}{P(M)} = \frac{\frac{2}{3} \cdot \frac{1}{3}}{\frac{1}{2}} = \frac{4}{9}$$

$$P(NBA|F) = \frac{2}{9}$$

* Finally, we'll be using the Poisson distribution in this class

def A Poisson(λ) Random Variable with $\lambda > 0$ can take on any nonnegative integer value $x = 0, 1, 2, \dots$ and has probability mass function

$$P(X=x) = \frac{\lambda^x}{x!} e^{-\lambda}$$

$$E X = \sum_x x P(X=x)$$

$$= \sum_{x=0}^{\infty} x e^{-\lambda} \frac{\lambda^x}{x!}$$

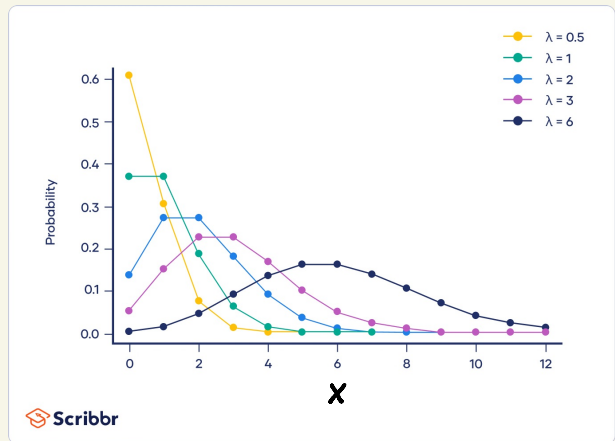
$$= \lambda e^{-\lambda} \sum_{x=1}^{\infty} \frac{\lambda^{x-1}}{(x-1)!}$$

$$= \lambda e^{-\lambda} \sum_{x=0}^{\infty} \frac{\lambda^x}{x!}$$

$$= \lambda e^{-\lambda} e^{\lambda} \text{ by def of Taylor series of } e^{\lambda}$$

$$= \lambda.$$

Similarly, $\text{var}(X) = \lambda.$



There are a few other "canonical" distributions you should know

- Geometric
- Hypergeometric
- Exponential
- Beta
- Gamma

HW Watch Professor Wyner's probability lectures from Moneyball (e.g., expected value)