

The Power of Fake Data (Priors)

Q Suppose the Dodgers have won W and lost L games thus far in the season.

How would you predict their end of season win percentage WP ?

- 162 total games in the season
- no access to their schedule (e.g., ignore strength of schedule)
- Without using previous season's data (i.e., Regression).

Guess their end of season win percentage.

Naive guess (ask any rando on the street):

$$\widehat{WP} = \frac{W}{W+L}$$

What's wrong with this?

When Dodgers have only played a few games, this estimate is bad.

Ex $W=3, L=0, \widehat{WP}=1$

Idea Add fake data.

Suppose the Dodgers begin the season with W' wins and L' losses.

New guess:

$$\widehat{WP}' = \frac{W+W'}{W+W'+L+L'}$$

For concreteness:

$$W=3, L=0, \widehat{WP}=1$$

Tom Tango: $W'=L'=15$ is good

$$W=3, L=0, W'=15, L'=15, \widehat{WP}' = \frac{18}{33} \approx .55$$

quite different predictions early in the season

$$W=45, L=30, \widehat{WP} = \frac{45}{75} = .6$$

$$W=45, L=30, W'=15, L'=15, \widehat{WP}' = \frac{60}{105} \approx .67$$

similar predictions late in the season

Which is better?

Formalize this

Dodgers play $n=162$ games in a season.

Suppose, for simplicity, that the Dodgers win each game with probability p .

Game outcomes $\{X_1, \dots, X_n\}$, where

$$X_i \sim \begin{cases} 1 & \text{w.p. } p \\ 0 & \text{w.p. } 1-p \end{cases} \stackrel{d}{=} \text{Bernoulli}(p)$$

Suppose we have observed m games thus far in the season.

Observed data $\{X_1, \dots, X_m\}$. Each X_i is 1 or 0.

Observed # wins $W = \sum_{i=1}^m X_i$.

So, $W \sim \text{Binomial}(m, p)$

$m = \# \text{ trials (games)}$
 $p = \text{prob. success (win)}$

and end-of-season win percentage $WP \sim \frac{1}{n} \text{Binomial}(n, p)$

Idea. Use observed data to estimate p , call it \hat{p}

Then, estimate $\widehat{WP} = \frac{1}{n} \mathbb{E}[\text{Binomial}(n, \hat{p})] = \frac{1}{n} \cdot n\hat{p} = \hat{p}$.

Maximum Likelihood estimate (MLE)

Choose \hat{p} to be the value of p which maximizes the probability of observing the game outcomes $\{x_1, \dots, x_m\}$ that we observed.

$$\hat{p}_{MLE} = \operatorname{argmax}_p \underbrace{P(x_1, \dots, x_m | p)}$$

likelihood: $P(\text{data given parameter})$

$$= \operatorname{argmax}_p P(x_1 | p) \cdot P(x_2 | p) \cdot \dots \cdot P(x_m | p)$$

by independence

$$= \operatorname{argmax}_p \prod_{i=1}^m P(x_i | p)$$

by def of product

$$= \operatorname{argmax}_p \prod_{i=1}^m p^{x_i} (1-p)^{1-x_i}$$

because $x_i \sim \text{BER}(p)$

$$x_i = 1 \text{ means } p^{x_i} (1-p)^{1-x_i} = p$$

$$x_i = 0 \text{ means } p^{x_i} (1-p)^{1-x_i} = 1-p$$

$$= \operatorname{argmax}_p p^{\sum_{i=1}^m x_i} (1-p)^{\sum_{i=1}^m (1-x_i)}$$

$$= \operatorname{argmax}_p p^W (1-p)^L$$

where $W = \sum_{i=1}^m x_i =$ number of wins (ones)

$L = \sum_{i=1}^m (1-x_i) =$ number of losses (zeros)

$$= \operatorname{argmax}_p \log [p^W \cdot (1-p)^L]$$

because \log is monotonic increasing

to maximize $f(p)$ it to maximize $\log f(p)$

$$= \operatorname{argmax}_p W \log p + L \log (1-p)$$

to maximize the function $p \mapsto W \log p + L \log(1-p)$
take the derivative and set it equal to 0
(and check that the 2nd derivative is negative).

$$\frac{d}{dp} [W \log p + L \log(1-p)]$$

$$= W \cdot \frac{1}{p} - L \cdot \frac{1}{1-p} = 0$$

$$\Rightarrow \frac{W}{p} = \frac{L}{1-p} \Rightarrow p = \frac{W}{L} (1-p)$$

$$\Rightarrow p \left(1 + \frac{W}{L}\right) = \frac{W}{L} \Rightarrow p = \frac{\frac{W}{L}}{1 + \frac{W}{L}}$$

$$\Rightarrow \hat{p}_{MLE} = \frac{W}{W+L} \quad \text{same formula from earlier!!}$$

The MLE is simply the observed win percentage midway through the season!

But we know this is a bad estimate early in the season.

So, why did the MLE go wrong??

How do we add the fake data W', L' to the MLE to get $\frac{W+W'}{W+W'+L+L'}$??

Before, to improve our estimate of WP , we added some fake data (W' , L').

In adding fake data, we used **prior information**:

Prior to the season, we assumed the Phillies have W' wins and L' losses.

What is a way of formalizing prior information?

Bayesian statistics — the belief/philosophy that we should treat a parameter (e.g. p) as having a probability distribution

Frequentist statistics — treats a parameter as an unknown fixed number

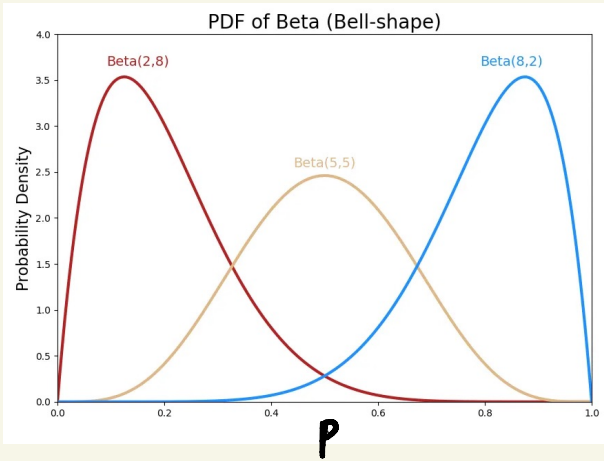
So, our way of formalizing the addition of prior "fake" data is to, prior to seeing the data, give a probability distribution to the parameter (e.g. p) which reflects our prior belief on what p is more likely to be than not!

Formally, we use the Beta-Binomial model :

$$\begin{cases} W \sim \text{Binomial}(m, P) \\ P \sim \text{Beta}(\alpha, \beta) \rightarrow \text{Prior} \\ \alpha = W' + 1, \quad \beta = L' + 1 \end{cases}$$

Beta distribution has density $f(p|\alpha, \beta) = C \cdot P^{\alpha-1} (1-P)^{\beta-1}$

on the interval $p \in [0, 1]$, where C is a constant chosen so that the distribution integrates to 1.



For example, $P \sim \text{Beta}(5,5)$ encodes a preference that P is closer to 0.5

As before, we wish to estimate p , this time with a Maximum a-Posteriori (MAP) Estimate:

Choose the \hat{p} which maximizes the posterior probability of p .

Bayesian Approach
to Parameter
Estimation

1. Prior
2. observe data
3. adjust our posterior dist for p given the data

$$\hat{p}_{\text{MAP}} = \underset{p}{\operatorname{argmax}} \underbrace{P(p|w)}_{\text{posterior} = P(\text{parameter} | \text{data})}$$

$$= \underset{p}{\operatorname{argmax}} \frac{P(w|p) \cdot P(p)}{P(w)} \quad \text{by Bayes' Rule}$$

$$= \underset{p}{\operatorname{argmax}} \underbrace{P(w|p)}_{\text{likelihood}} \cdot \underbrace{P(p)}_{\text{prior}}$$

since $P(w)$ has no p term

$$= \underset{p}{\operatorname{argmax}} P(\text{Binomial}(n, p) = w) \cdot P(\text{beta}(a, b) = p)$$

$$= \operatorname{argmax}_p \binom{m}{w} p^w (1-p)^{m-w} \cdot C p^{\alpha-1} (1-p)^{\beta-1}$$

$$= \operatorname{argmax}_p p^w (1-p)^L \cdot p^{\alpha-1} (1-p)^{\beta-1}$$

$$= \operatorname{argmax}_p p^{w+\alpha-1} (1-p)^{L+\beta-1}$$

= ... same process as before

$$= \frac{w+\alpha-1}{w+\alpha-1 + L+\beta-1}$$

$$= \frac{w+w'}{w+w'+L+L'} \quad \text{if} \quad \begin{array}{l} w' = \alpha-1 \\ L' = \beta-1 \end{array}$$

The MAP estimate is simply the win percentage if we add $\alpha-1$ fake wins and $\beta-1$ fake losses!!

Can use past seasons to tune a smart choice for α, β .

Note: $\alpha=1, \beta=1 \Rightarrow \hat{p}_{\text{MAP}} = \hat{p}_{\text{MLE}}$
add no fake data

Model $\begin{cases} W \sim \text{Binomial}(n, p) \\ p \sim \text{Uniform}(0, 1) \end{cases} \rightarrow$ uninformative
prior which encodes
no preference on p

$$\begin{aligned} \hat{p}^{(\text{MAP})} &= \underset{p}{\text{argmax}} P(p|W) = \underset{p}{\text{argmax}} P(W|p) \cdot \underbrace{P(p)}_1 \\ &= \underset{p}{\text{argmax}} P(W|p) = \hat{p}^{(\text{MLE})} = \frac{W}{W+L} \end{aligned}$$

Takeaways

- Bayesian Statistics: treat a parameter (e.g., p) as having a distribution
- Blend observed data with prior knowledge, encoding info not seen in the data, to make better predictions